# General Cardiovascular Risk Profile for Use in Primary Care The Framingham Heart Study 

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Background-Separate multivariable risk algorithms are commonly used to assess risk of specific atherosclerotic cardiovascular disease (CVD) events, ie, coronary heart disease, cerebrovascular disease, peripheral vascular disease, and heart failure. The present report presents a single multivariable risk function that predicts risk of developing all CVD and of its constituents.
Methods and Results-We used Cox proportional-hazards regression to evaluate the risk of developing a first CVD event in 8491 Framingham study participants (mean age, 49 years; 4522 women) who attended a routine examination between 30 and 74 years of age and were free of CVD. Sex-specific multivariable risk functions ("general CVD" algorithms) were derived that incorporated age, total and high-density lipoprotein cholesterol, systolic blood pressure, treatment for hypertension, smoking, and diabetes status. We assessed the performance of the general CVD algorithms for predicting individual CVD events (coronary heart disease, stroke, peripheral artery disease, or heart failure). Over 12 years of follow-up, 1174 participants ( 456 women) developed a first CVD event. All traditional risk factors evaluated predicted CVD risk (multivariable-adjusted $P<0.0001$ ). The general CVD algorithm demonstrated good discrimination (C statistic, 0.763 [men] and 0.793 [women]) and calibration. Simple adjustments to the general CVD risk algorithms allowed estimation of the risks of each CVD component. Two simple risk scores are presented, 1 based on all traditional risk factors and the other based on non-laboratory-based predictors.
Conclusions-A sex-specific multivariable risk factor algorithm can be conveniently used to assess general CVD risk and risk of individual CVD events (coronary, cerebrovascular, and peripheral arterial disease and heart failure). The estimated absolute CVD event rates can be used to quantify risk and to guide preventive care. (Circulation. 2008;117: 743-753.)

Key Words: cardiovascular diseases ■ coronary disease ■ heart failure ■ risk factors ■ stroke

It is widely accepted that age, sex, high blood pressure, smoking, dyslipidemia, and diabetes are the major risk factors for developing cardiovascular disease (CVD). ${ }^{1}$ It also is recognized that CVD risk factors cluster and interact multiplicatively to promote vascular risk. ${ }^{2}$ This knowledge led to the development of multivariable risk prediction algorithms incorporating these risk factors that can be used by primary care physicians to assess in individual patients the risk of developing all atherosclerotic $\mathrm{CVD}^{3-12}$ or specific components of CVD, ie, coronary heart disease, ,,13-17 stroke, ${ }^{18}$ peripheral vascular disease, ${ }^{19}$ or heart failure. ${ }^{20}$ Multivariable assessment has been advocated to estimate absolute CVD risk and to guide treatment of risk factors. ${ }^{2,6}$ For instance, the Framingham formulation for predicting coronary heart disease (CHD) was incorporated into the Third

Report of the Expert Panel on Detection, Evaluation, and Treatment of High Blood Cholesterol in Adults (Adult Treatment Panel III). ${ }^{9}$ The Framingham CHD risk assessment tool has been validated in whites and blacks in the United States ${ }^{9,10,21}$ and are transportable (with calibration) to culturally diverse populations in Europe, the Mediterranean region, and Asia. ${ }^{9,10,22,23}$ Similar CHD risk prediction algorithms have been developed by other investigators worldwide and have been demonstrated to perform well. ${ }^{14,15,17}$

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Despite the availability of several validated risk prediction algorithms, their use has lagged in primary care. ${ }^{24}$ One potential reason for physician inertia in using risk prediction instruments is the multiplicity of such algorithms, each for

[^0]Circulation is available at http://circ.ahajournals.org
predicting an individual CVD component. Indeed, there are occasions when a physician would like to target risk assessment and preventive measures to a specific cardiovascular end point such as myocardial infarction or stroke depending, for example, on an individual patient's family history, age, diabetic status, or predisposition to a particular outcome by valve disease. However, with this exception, primary care physicians engaged in preventive health maintenance want to assess risk of developing any major atherosclerotic CVD event using a general CVD risk assessment tool. Accordingly, the purpose of the present investigation was to formulate a single multivariable risk assessment tool that would enable physicians to identify high-risk candidates for any and all initial atherosclerotic CVD events using measurements readily available at the clinic or office.

## Methods

## Study Design and Sample

The design and selection criteria for the original Framingham Heart Study and the Framingham Offspring Study have been detailed elsewhere. ${ }^{25,26}$ Detailed descriptions of the examination procedures and criteria for CVD events also have been reported. ${ }^{27}$ Participants were eligible for the present investigation if they attended the 11th biennial examination cycle of original cohort (1968 to 1971, when measurement of high-density lipoprotein [HDL] cholesterol was available) or the first (1971 to 1975) or third (1984 to 1987) examination cycles of the Offspring cohort and were free of CVD. All participants provided written informed consent, and the study protocol was approved by the Institutional Review Board at the Boston Medical Center.

The study sample consisted of attendees of the baseline examinations free of prevalent CVD who were 30 to 74 years of age with nonmissing data on covariates. After exclusions, 8491 participants (mean age, 49 years; 4522 women) remained eligible.

## Measurement of CVD Risk Factors

At each heart study examination, participants underwent a physical examination, anthropometry, blood pressure determination, and phlebotomy for vascular risk factors. Blood pressure measurements were made on the left arm of the seated participants with a mercury-column sphygmomanometer and an appropriately sized cuff; the average of 2 physician-obtained measures constituted the examination blood pressure. Serum total and HDL cholesterol levels were determined with standardized enzymatic methods. Cigarette smoking status was ascertained by self-report. Diabetes was defined as fasting glucose $\geq 126 \mathrm{mg} / \mathrm{dL}$ (offspring cohort) or $140 \mathrm{mg} / \mathrm{dL}$ (original cohort) or use of insulin or oral hypoglycemic medications. Antihypertensive medication use was ascertained by the physician examiner at the heart study and based on self-report.

## Follow-Up and Outcome Events

All study participants were under continuous surveillance for the development of CVD events and death. The Framingham Heart Study defines CVD as a composite of CHD (coronary death, myocardial infarction, coronary insufficiency, and angina), cerebrovascular events (including ischemic stroke, hemorrhagic stoke, and transient ischemic attack), peripheral artery disease (intermittent claudication), and heart failure. ${ }^{1}$ Information about CVD events on follow-up was obtained with the aid of medical histories, physical examinations at the study clinic, hospitalization records, and communication with personal physicians. All suspected new events were reviewed by a panel of 3 experienced investigators who evaluated all pertinent medical records. A separate review committee that included a neurologist adjudicated cerebrovascular events, and a heart study neurologist examined most participants with suspected stroke.

## Statistical Analyses

## Multivariable Models and Estimation of General CVD Risk Functions

We used sex-specific Cox proportional-hazards regressions ${ }^{28}$ to relate risk factors to the incidence of a first CVD event during a maximum follow-up period of 12 years after confirming that the assumption of proportionality of hazards was met. From these models, we estimated mathematical CVD risk functions, ${ }^{28}$ referred to as a general CVD risk function (Appendix); these functions were used to estimate 10 -year absolute CVD risk.

Covariates included in Cox models were age, total cholesterol, HDL cholesterol, systolic blood pressure, antihypertensive medication use, current smoking, and diabetes status. Other variables such as diastolic blood pressure, body mass index, and triglycerides also were considered, but they were not statistically significant. The use of low-density lipoprotein cholesterol did not improve model fit or performance. All the continuous variables were naturally logarithmically transformed to improve discrimination and calibration of the models and to minimize the influence of extreme observations. We adjusted for the use of antihypertensive medication by modeling the impact of a participant's systolic blood pressure differently on the basis of use of such medications.

## Assessment of Model Performance

We evaluated the ability of the risk prediction model to discriminate persons who experience a CVD event from those who do not using an overall c statistic, ${ }^{29,30}$ expanding on a suggestion by Harrell et al. ${ }^{31}$ This c statistic is analogous to the area under the receiveroperating characteristic curve. Briefly, 2 subjects are described as comparable if we can determine which one survived longer and concordant if their predicted probabilities of survival and survival times go in the same direction, and we can define the overall c statistic as the probability of concordance given comparability. The degree of overoptimism resulting from model assessment on the same data on which it was developed was estimated on the basis of bootstrap resampling of the original set.

We evaluated the calibration of our risk prediction model, a measure of agreement between observed and predicted events within 10 years, using a modified Hosmer-Lemeshow $\chi^{2}$ statistic with 9 $d f .{ }^{29}$ For this purpose, we used the Kaplan-Meier estimator to obtain the observed incidence of CVD events, which was then compared with the CVD risk predicted by the model and classified into deciles. ${ }^{29}$ We also calculated the proportion of CVD events that occurred in the top quintile of predicted risk (ie, sensitivity of the top quintile of predicted risk for identifying CVD events) and the proportion of individuals without events who are not in the top quintile of predicted risk (ie, specificity of the top quintile for CVD events).
The performance of the new CVD risk prediction model presented here was compared with that of another popular Framingham risk score developed by Wilson et al. ${ }^{16}$ Because the latter score was developed for predicting CHD and not CVD, we performed a simple recalibration by multiplying the risk of each individual by the ratio of CVD incidence rate and the mean predicted risk based on the CHD risk function. Thus, we assessed how well the Framingham CHD risk functions ${ }^{16}$ predicted CVD relative to the new CVD prediction model. A test for difference in 2 correlated c statistics proposed by Antolini et al ${ }^{32}$ was used, along with the net reclassification improvement proposed by Pencina et al. ${ }^{33}$ Reclassification improvement is defined as an increase in risk category for individuals who develop events and as a decrease for those who do not. Net reclassification improvement accounts for movement between categories in the wrong direction and applies different weights to events and nonevents. We used $0 \%$ to $6 \%, 6 \%$ to $20 \%$, and $>20 \%$ as risk categories.

## Performance of General CVD Risk Prediction Model for Predicting Individual CVD Components

After generating sex-specific general CVD risk functions as detailed above, we applied them to predict the risk of individual components

Table 1. Summary Statistics for Risk Factors Used in Risk Models

|  | Women | Men |
| :--- | :---: | :---: |
| Characteristics | $(\mathrm{n}=4522,28 \%$ FOC $)$ | $(\mathrm{n}=3969,22 \% \mathrm{FOC})$ |
| Age, mean (SD), y | $49.1(11.1)$ | $48.5(10.8)$ |
| Total-C, mean (SD), mg/dL | $215.1(44.1)$ | $212.5(39.3)$ |
| HDL-C, mean (SD), mg/dL | $57.6(15.3)$ | $44.9(12.2)$ |
| Systolic BP, mean (SD), | $125.8(20.0)$ | $129.7(17.6)$ |
| mm Hg |  |  |
| BP treatment, $\mathrm{n}(\%)$ | $532(11.76)$ | $402(10.13)$ |
| Smoking, $\mathrm{n}(\%)$ | $1548(34.23)$ | $1398(35.22)$ |
| Diabetes, $\mathrm{n}(\%)$ | $170(3.76)$ | $258(6.50)$ |
| Incident CVD events, $\mathrm{n}(\%)$ | $456(10.08)$ | $718(18.09)$ |

FOC indicates Framingham original cohort; Total-C, total cholesterol; HDL-C, HDL cholesterol; and BP, blood pressure.
of CVD (CHD, stroke, intermittent claudication, congestive heart failure) after multiplication of the probability predicted by the general risk function by the proportion of all CVD events that were constituted by an individual component (ratio of Kaplan-Meier event rates). These were contrasted with models that we developed for individual CVD components using the same predictors.

## Sex-Specific General CVD Risk Scores Sheets

## and Heart Age

General CVD risk functions were translated into sex-specific risk score sheets by use of previously described methods. ${ }^{34}$ To facilitate easier understanding of the concept of risk, we also constructed "heart age" sheets. An individual's heart age is calculated as the age of a person with the same predicted risk but with all other risk factor levels in normal ranges. Although called heart age for simplicity of risk communication in primary care, the heart age really reflects vascular age. In the following, we use heart age/vascular age.

## Simpler CVD Risk Prediction Models Using Nonlaboratory Predictors Routinely Ascertained in Primary Care

In addition to the main CVD risk prediction models described above, we developed simplified sex-specific models that used simple
office-based predictors that are routinely obtained in primary care and do not require laboratory testing. These variables included age, body mass index, systolic blood pressure, antihypertensive medication use, current smoking, and diabetes status. The same modeling principles and model assessment techniques were applied to these simplified models.
The authors had full access to and take full responsibility for the integrity of the data. All authors have read and agree to the manuscript as written.

## Results

The risk factor characteristics of men and women in our sample at the baseline examinations are shown in Table 1. In our middle-aged sample, mean levels of serum total cholesterol and systolic blood pressure were similar in men and women, as were the prevalences of cigarette smoking and use of antihypertensive treatment. The prevalence of diabetes was substantially higher in men, whereas mean serum HDL levels were higher in women.

## General CVD Risk Prediction Models

The multivariable-adjusted regression coefficients and hazard ratios for incident CVD events are presented in Table 2. We observed highly statistically significant relations of all risk factors evaluated and incident CVD.

The sex-specific CVD functions performed well in terms of both model discrimination and calibration. The c statistics for the risk function ranged from 0.763 ( $95 \%$ confidence interval [CI], 0.746 to 0.780 ) in men to 0.793 ( $95 \%$ CI, 0.772 to 0.814 ) in women. The degree of overoptimism was estimated at 0.001 for men and 0.003 for women, partly reflecting a large number of events and the potential limitation of the bootstrap resampling approach for assessing overoptimism.

The calibration $\chi^{2}$ statistics for the CVD prediction models were 13.48 in men and 7.79 for the women, indicating excellent goodness of fit (for the lack of fit, $P=0.14$ and

Table 2. Regression Coefficients and Hazard Ratios

| Variable | $\beta^{\star}$ | P | Hazard Ratio | $95 \% \mathrm{Cl}$ |
| :--- | ---: | :--- | ---: | :--- |
| Women [So(10)=0.95012] |  |  |  |  |
| Log of age | 2.32888 | $<0.0001$ | 10.27 | $(5.65-18.64)$ |
| Log of total cholesterol | 1.20904 | $<0.0001$ | 3.35 | $(2.00-5.62)$ |
| Log of HDL cholesterol | -0.70833 | $<0.0001$ | 0.49 | $(0.35-0.69)$ |
| Log of SBP if not treated | 2.76157 | $<0.0001$ | 15.82 | $(7.86-31.87)$ |
| Log of SBP if treated | 2.82263 | $<0.0001$ | 16.82 | $(8.46-33.46)$ |
| Smoking | 0.52873 | $<0.0001$ | 1.70 | $(1.40-2.06)$ |
| $\quad$ Diabetes | 0.69154 | $<0.0001$ | 2.00 | $(1.49-2.67)$ |
| Men [So(10)=0.88936] |  |  |  |  |
| Log of age | 3.06117 | $<0.0001$ | 21.35 | $(14.03-32.48)$ |
| Log of total cholesterol | 1.12370 | $<0.0001$ | 3.08 | $(2.05-4.62)$ |
| Log of HDL cholesterol | -0.93263 | $<0.0001$ | 0.39 | $(0.30-0.52)$ |
| Log of SBP if not treated | 1.93303 | $<0.0001$ | 6.91 | $(3.91-12.20)$ |
| Log of SBP if treated | 1.99881 | $<0.0001$ | 7.38 | $(4.22-12.92)$ |
| Smoking | 0.65451 | $<0.0001$ | 1.92 | $(1.65-2.24)$ |
| Diabetes | 0.57367 | $<0.0001$ | 1.78 | $(1.43-2.20)$ |

So(10) indicates 10-year baseline survival; SBP, systolic blood pressure.
*Estimated regression coefficient


Figure. Calibration by decile for CVD function for women (A) and men (B). Vertical bars represent observed (Kaplan-Meier [km]; black) and model-based predicted (decile specific means; gray) probabilities of CVD event in 10 years in deciles of modelbased predicted probabilities.
$P=0.56$, respectively). The Figure displays the calibration plots comparing predicted deciles of risk and actual observed risk in men and women. The top sex-specific quintiles of predicted risk identified $\approx 49 \%$ of men and $60 \%$ of women who experienced a first CVD event on follow-up (sensitivity). Proportions of men and women without CVD events who were not in the top quintile of predicted risk were $85 \%$ and $84 \%$, respectively (specificity).

The Framingham CHD risk functions (Wilson et al ${ }^{16}$ ) performed less well for predicting CVD risk: The c statistics were lower ( 0.756 [ $95 \% \mathrm{CI}, 0.739,0.773$ ] in men; for difference compared with our new model, $P=0.051 ; 0.778$ [ $95 \%$ CI, 0.756, 0.799] in women; for difference compared with our new model, $P=0.003$ ) and calibration was worse ( $\chi^{2}=32.37$ in men and 12.42 in women) relative to that noted above for the new CVD risk prediction models. The sensitivity of the top quintile of predicted risk using the CHD risk functions was slightly lower ( $47 \%$ in men and $56 \%$ in women) although specificity was similar ( $85 \%$ in men and $83 \%$ in women). The net reclassification improvement from using the new model was statistically significant for both men and women and reached $6.65 \%(P<0.001)$ and $7.95 \%$ ( $P=0.003$ ), respectively.

## Performance of General CVD Risk Prediction <br> Model for Predicting Individual CVD Components

Tables 3 and 4 assess the performances of the sex-specific general CVD risk functions by comparing them with disease-

Table 3. Performance Summary: Modified CVD Model Versus Event-Specific Own Model for Women

|  | CVD Model | Own Model |
| :---: | :---: | :---: |
| CHD ( $\mathrm{n}=216$ ) |  |  |
| C | 0.787 | 0.789 |
| 95\% Cl for C | (0.762-0.812) | (0.764-0.815) |
| $\chi^{2}$ | 14.79 | 17.52 |
| $P$ for $\chi^{2}$ | 0.097 | 0.041 |
| Sensitivity of top quintile | 57.55 | 56.38 |
| Specificity of top quintile | 81.94 | 81.88 |
| Calibration factor | 0.6086 |  |
| So(10) |  | 0.9704 |
| Stroke ( $\mathrm{n}=84$ ) |  |  |
| C | 0.769 | 0.774 |
| 95\% CI for C | (0.715-0.822) | (0.721-0.828) |
| $\chi^{2}$ | 5.26 | 6.86 |
| $P$ for $\chi^{2}$ | 0.811 | 0.651 |
| Sensitivity of top quintile | 61.56 | 63.91 |
| Specificity of top quintile | 80.82 | 80.86 |
| Calibration Factor | 0.2385 |  |
| So(10) |  | 0.9898 |
| CHF ( $\mathrm{n}=44$ ) |  |  |
| C | 0.847 | 0.851 |
| 95\% Cl for C | (0.803-0.891) | (0.804-0.897) |
| $\chi^{2}$ | 9.32 | 8.82 |
| $P$ for $\chi^{2}$ | 0.408 | 0.454 |
| Sensitivity of top quintile | 76.49 | 83.73 |
| Specificity of top quintile | 80.58 | 80.65 |
| Calibration factor | 0.1250 |  |
| So(10) |  | 0.9962 |
| IC ( $\mathrm{n}=66$ ) |  |  |
| C | 0.829 | 0.848 |
| 95\% CI for C | (0.786-0.872) | (0.810-0.887) |
| $\chi^{2}$ | 11.33 | 11.63 |
| $P$ for $\chi^{2}$ | 0.254 | 0.235 |
| Sensitivity of top quintile | 70.25 | 70.07 |
| Specificity of top quintile | 80.77 | 80.76 |
| Calibration factor | 0.1862 |  |
| So(10) |  | 0.9918 |

C indicates model discrimination (c statistic); Sensitivity of top quintile, percent events captured by the top quintile of predicted risk; Specificity of top quintile, percent nonevents captured by the bottom 4 quintiles of predicted risk; So(10), baseline survival rate at 10 years; and IC, intermittent claudication.
specific algorithms for predicting risk of CHD, stroke, intermittent claudication, and heart failure. To apply the CVD functions for a specific component, the CVD-predicted probabilities were multiplied by the "calibration factor" given in Tables 3 and 4 . For example, to compute the 10 -year probability of CHD from the general CVD risk function in women, the CVD probability is calculated and then multiplied by 0.61 , the proportion of first CVD events in women that were CHD events.

Table 4. Performance Summary: Modified CVD Model Versus Event-Specific Own Model for Men

|  | CVD Model | Own Model |
| :---: | :---: | :---: |
| CHD ( $\mathrm{n}=425$ ) |  |  |
| C | 0.733 | 0.735 |
| 95\% Cl for C | (0.712-0.754) | (0.714-0.756) |
| $\chi^{2}$ | 18.20 | 18.36 |
| $P$ for $\chi^{2}$ | 0.033 | 0.031 |
| Sensitivity of top quintile | 45.94 | 45.70 |
| Specificity of top quintile | 83.23 | 83.20 |
| Calibration factor | 0.7174 |  |
| So(10) |  | 0.9167 |
| Stroke ( $\mathrm{n}=93$ ) |  |  |
| C | 0.826 | 0.835 |
| 95\% Cl for C | (0.789-0.863) | (0.797-0.874) |
| $\chi^{2}$ | 26.11 | 9.21 |
| $P$ for $\chi^{2}$ | 0.002 | 0.418 |
| Sensitivity of top quintile | 71.64 | 76.05 |
| Specificity of top quintile | 81.30 | 81.41 |
| Calibration factor | 0.1590 |  |
| So(10) |  | 0.9883 |
| CHF ( $\mathrm{n}=67$ ) |  |  |
| C | 0.841 | 0.845 |
| 95\% Cl for C | (0.799-0.883) | (0.802-0.888) |
| $\chi^{2}$ | 27.23 | 15.30 |
| $P$ for $\chi^{2}$ | 0.001 | 0.083 |
| Sensitivity of top quintile | 80.55 | 82.59 |
| Specificity of top quintile | 81.09 | 81.13 |
| Calibration factor | 0.1148 |  |
| So(10) |  | 0.9927 |
| IC ( $\mathrm{n}=105$ ) |  |  |
| C | 0.813 | 0.820 |
| 95\% Cl for C | (0.780-0.847) | (0.787-0.853) |
| $\chi^{2}$ | 19.05 | 8.18 |
| $P$ for $\chi^{2}$ | 0.025 | 0.516 |
| Sensitivity of top quintile | 60.29 | 66.65 |
| Specificity of top quintile | 81.15 | 81.34 |
| Calibration factor | 0.1804 |  |
| So(10) |  | 0.9852 |

Abbreviations as in Table 3.

From the comparison of discrimination and $\chi^{2}$ statistics, it is evident that the general CVD risk formulation provides discrimination of individual CVD outcomes that is as good as the individual disease-specific multivariable risk formulations and is well calibrated. Similarly, the sensitivity of the upper quintile of the CVD risk function is comparable to that of the top quintile of disease-specific functions in both sexes (Tables 3 and 4). In the analyses of individual components, the regression coefficients for cholesterol were higher for CHD and intermittent claudication (relative to that for stroke and congestive heart failure [CHF]; data not shown). Systolic blood pressure was more strongly associated with stroke and CHF (compared with CHD and intermittent claudication),
and smoking was more strongly associated with intermittent claudication (data not shown).

## Derivation of CVD Prediction Scores and Heart Age/Vascular Age

Tables 5 and 6 and 7 and 8 provide score sheets that can be used for estimating the multivariable risk of CVD for women and men, respectively. Tables 9 and 10 give a different quantification of the same risk in the form of heart age/vascular age. We illustrate the use of these tables in the Appendix, and they are available at www.framinghamheartstudy.org/risk/index.html.

## Simpler CVD Risk Prediction Models Using Nonlaboratory Predictors

The simple office-based CVD risk prediction function that incorporated body mass index (instead of total and HDL cholesterol) performed reasonably well (Table I of the online Data Supplement). The discrimination c statistics was 0.749 ( $95 \%$ CI, $0.731,0.767$ ) for men and 0.785 ( $95 \%$ CI, 0.764 , 0.806 ) for women (for difference compared with our full model, $P<0.001$ and 0.013 , respectively). Calibration $\chi^{2}$ statistics were 13.61 (for the lack of fit, $P=0.14$ ) for men and 10.24 for women (for the lack of fit, $P=0.33$ ). The top sex-specific quintiles of predicted CVD risk identified $\approx 48 \%$ of men and $58 \%$ of women who experienced a first CVD event on follow-up (sensitivity). Proportions of men and women without events who were not in the top quintile of risk were $85 \%$ and $83 \%$, respectively (specificity). Tables IIA through IIC and IIIA through IIIC in the Data Supplement provide score sheets that can be used to estimate the multivariable risk of CVD and heart age/vascular age for women and men, respectively, using the office-based nonlaboratory predictors.

## Discussion

It is widely accepted that CVD constitutes a major public health problem in the United States ${ }^{35}$ and worldwide. ${ }^{36}$ The lifetime risk of CVD is substantial, ${ }^{37}$ and the condition is often silent or may strike without warning, underscoring the importance of prevention. Investigators have identified key risk factors that account for most CVD burden in the community, and numerous reports have demonstrated the clustering and conjoint influences of multiple risk factors in mediating disease vascular risk. ${ }^{2,4,6,7,38-41}$ Consequently, researchers have devised multivariable risk prediction tools that synthesize vascular risk factor information to yield estimates of absolute CVD risk (also referred to as global CVD risk) in individual patients. $4,7,8,10-12,42$ The estimation of global CVD risk facilitates the matching of the intensity of risk factor lowering with the estimated probability of disease, thereby rendering treatment most cost-effective. ${ }^{38,42-44}$ For instance, national cholesterol guidelines link treatment thresholds and goals to global coronary heart disease risk. ${ }^{9}$ In addition to reducing the number needed to treat to prevent a CVD event, multivariable risk assessment also avoids overlooking highrisk CVD candidates with multiple marginal risk factors and avoids needlessly alarming persons with only 1 isolated risk factor. Furthermore, analyses that fail to examine risk factors

Table 5. CVD Points for Women

| Points | Age, y | HDL | Total Cholesterol | SBP Not Treated | SBP Treated | Smoker | Diabetic |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -3 |  |  |  | $<120$ |  |  |  |
| -2 |  | $60+$ |  |  |  |  |  |
| -1 |  | 50-59 |  |  | $<120$ |  |  |
| 0 | 30-34 | 45-49 | <160 | 120-129 |  | No | No |
| 1 |  | 35-44 | 160-199 | 130-139 |  |  |  |
| 2 | 35-39 | <35 |  | 140-149 | 120-129 |  |  |
| 3 |  |  | 200-239 |  | 130-139 | Yes |  |
| 4 | 40-44 |  | 240-279 | 150-159 |  |  | Yes |
| 5 | 45-49 |  | $280+$ | 160+ | 140-149 |  |  |
| 6 |  |  |  |  | 150-159 |  |  |
| 7 | 50-54 |  |  |  | 160+ |  |  |
| 8 | 55-59 |  |  |  |  |  |  |
| 9 | 60-64 |  |  |  |  |  |  |
| 10 | 65-69 |  |  |  |  |  |  |
| 11 | 70-74 |  |  |  |  |  |  |
| 12 | 75+ |  |  |  |  |  |  |
| Points allotted |  |  |  |  |  |  |  |

SBP indicates systolic blood pressure.
in combinations usually greatly overestimate the populationattributable risks associated with individual risk factors. ${ }^{45}$

Researchers also have developed disease-specific formulations to predict risk of developing specific CVD events such

Table 6. CVD Risk for Women

| Points | Risk, $\%$ |
| ---: | :---: |
| $\leq-2$ | $<1$ |
| -1 | 1.0 |
| 0 | 1.2 |
| 1 | 1.5 |
| 2 | 1.7 |
| 3 | 2.0 |
| 4 | 2.4 |
| 5 | 2.8 |
| 6 | 3.3 |
| 7 | 3.9 |
| 8 | 4.5 |
| 9 | 5.3 |
| 10 | 6.3 |
| 11 | 7.3 |
| 12 | 8.6 |
| 13 | 10.0 |
| 14 | 11.7 |
| 15 | 13.7 |
| 16 | 15.9 |
| 17 | 18.5 |
| 18 | 21.5 |
| 19 | 24.8 |
| 20 | 30 |
|  |  |

as CHD events or stroke. ${ }^{13-16,18-20}$ The present investigation is based on the premise that although the impacts of risk factors vary from 1 specific CVD type to another, there is sufficient commonality of risk factors to warrant generating a single general CVD risk prediction instrument that could accurately predict global CVD risk and the risk of individual components. Our study was motivated by our presumption of a need to simplify risk prediction in office-based practices by replacing disease-specific algorithms with a single general CVD prediction tool.

Framingham investigators formulated a general CVD risk function several years ago. ${ }^{46}$ Using a multivariable-logistic regression model, we reported that an algorithm that identified persons at high risk of atherosclerotic CVD in general also was effective for identifying persons at risk for each of the specific events, including CHD, stroke, intermittent claudication, and heart failure. However, that risk formulation was developed in 1976; was based on a limited number of events; did not include HDL cholesterol, a powerful influence on lipid atherogenesis; and did not focus on estimates of absolute risk. A subsequent CVD risk function used a parametric model, but that investigation did not evaluate the ability of a general CVD risk profile to predict individual outcomes. ${ }^{3}$

The present investigation extends and expands on the previous general CVD risk formulation on the basis of a larger number of events, incorporates HDL cholesterol, and estimates absolute CVD risk. We propose a general CVD risk function that demonstrates very good discrimination and calibration both for predicting CVD and for predicting risk of individual CVD components (comparable to disease-specific algorithms). The parallelism between atherosclerosis in different vascular territories in terms of sharing a common set of risk factors explains why the general CVD risk function performs well for predicting the individual components. The

Table 7. CVD Points for Men

| Points | Age, y | HDL | Total Cholesterol | SBP Not Treated | SBP Treated | Smoker | Diabetic |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -2 |  | 60+ |  | <120 |  |  |  |  |
| -1 |  | 50-59 |  |  |  |  |  |  |
| 0 | 30-34 | 45-49 | <160 | 120-129 | <120 | No | No |  |
| 1 |  | 35-44 | 160-199 | 130-139 |  |  |  |  |
| 2 | 35-39 | <35 | 200-239 | 140-159 | 120-129 |  |  |  |
| 3 |  |  | 240-279 | $160+$ | 130-139 |  | Yes |  |
| 4 |  |  | $280+$ |  | 140-159 | Yes |  |  |
| 5 | 40-44 |  |  |  | $160+$ |  |  |  |
| 6 | 45-49 |  |  |  |  |  |  |  |
| 7 |  |  |  |  |  |  |  |  |
| 8 | 50-54 |  |  |  |  |  |  |  |
| 9 |  |  |  |  |  |  |  |  |
| 10 | 55-59 |  |  |  |  |  |  |  |
| 11 | 60-64 |  |  |  |  |  |  |  |
| 12 | 65-69 |  |  |  |  |  |  |  |
| 13 |  |  |  |  |  |  |  |  |
| 14 | 70-74 |  |  |  |  |  |  |  |
| 15 | 75+ |  |  |  |  |  |  |  |
| Points allotted |  |  |  |  |  |  |  | Total |

C statistic for the general CVD risk prediction models ranged from 0.76 to 0.79 , suggesting that additional risk factors may be considered in future studies for inclusion in the models to further improve model discrimination. The general CVD risk prediction function performed better than the Framingham

Table 8. CVD Risk for Men

| Points | Risk, $\%$ |
| :---: | :---: |
| $\leq-3$ or less | $<1$ |
| -2 | 1.1 |
| -1 | 1.4 |
| 0 | 1.6 |
| 1 | 1.9 |
| 2 | 2.3 |
| 3 | 2.8 |
| 4 | 3.3 |
| 5 | 3.9 |
| 6 | 4.7 |
| 7 | 5.6 |
| 8 | 6.7 |
| 9 | 7.9 |
| 10 | 9.4 |
| 11 | 11.2 |
| 12 | 13.2 |
| 13 | 15.6 |
| 14 | 18.4 |
| 15 | 21.6 |
| 16 | 25.3 |
| 17 | 29.4 |
| $18+$ | $>30$ |

CHD risk function ${ }^{16}$ for predicting CVD risk. The specific focus on CVD risk and the modeling of risk factors as continuous variables (as opposed to the use of categories in the CHD risk function developed by Wilson et $\mathrm{al}^{16}$ ) may explain the better performance of the former.

## Comparison With Other CVD

## Risk Prediction Tools

Although several instruments have been formulated to predict CHD, ${ }^{9,14,15,17}$ tools that predict CVD are fewer. For instance, the scoring system developed by the Prospective Cardiovascular Munster Study (PROCAM) investigators ${ }^{14}$ focuses on

Table 9. Heart Age/Vascular Age for Women

| Points | Heart Age, y |
| :---: | :---: |
| $<1$ | $<30$ |
| 1 | 31 |
| 2 | 34 |
| 3 | 36 |
| 4 | 39 |
| 5 | 42 |
| 6 | 45 |
| 7 | 48 |
| 8 | 51 |
| 9 | 55 |
| 10 | 59 |
| 11 | 64 |
| 12 | 68 |
| 13 | 73 |
| 14 | 79 |
| $15+$ |  |

Table 10. Heart Age/Vascular Age for Men

| Points | Heart Age, y |
| :---: | :---: |
| $<0$ | $<30$ |
| 0 | 30 |
| 1 | 32 |
| 2 | 34 |
| 3 | 36 |
| 4 | 38 |
| 5 | 40 |
| 6 | 42 |
| 7 | 45 |
| 8 | 48 |
| 9 | 51 |
| 10 | 54 |
| 11 | 57 |
| 12 | 60 |
| 13 | 64 |
| 14 | 68 |
| 15 | 72 |
| 16 | 76 |
| $\geq 17$ | $>80$ |

prediction of acute coronary events. The British ${ }^{4}$ and New Zealand ${ }^{7}$ guidelines use an older Framingham equation ${ }^{3}$ to facilitate prediction of global CVD risk. The present report offers an updated risk function based on a greater number of CVD events in a more contemporary time period with evaluation of calibration and "exchangeability" for diseasespecific risk profiles.

Ridker et al ${ }^{10}$ recently published a Reynolds risk score for predicting CVD in women. The Reynolds risk score incorporates family history of CVD, high-sensitivity C-reactive protein, and hemoglobin $\mathrm{A}_{1 \mathrm{c}}$ (the latter in individuals with diabetes). In addition, the Reynolds risk score was developed in women alone and did not include some CVD end points (such as intermittent claudication), and its transportability to other samples or its exchangeability for disease-specific profiles is unknown. It is conceivable that the general CVD risk algorithm proposed in the present investigation and the Reynolds risk score could be sequential components of a staged approach: The former is a simpler formulation using only lipids from several eligible candidate biomarkers, and the latter is a strategy that could be applied to a specific subgroup identified by the former that is targeted for measurement of additional biomarkers. The validity of such a premise of sequential testing warrants further research.

The Systematic Coronary Risk Evaluation (SCORE) project ${ }^{5}$ formulated a CVD risk estimation algorithm (HEARTSCORE) that has been adopted by the Joint European Societies' guidelines on CVD prevention. ${ }^{6}$ Whereas the SCORE risk functions have the advantage of being based on European epidemiological studies, the HEARTSCORE predicts only fatal CVD, which can result in an underestimation of the total CVD burden.

More recently, 2 CVD risk scores have been formulated by investigators in the United Kingdom. ${ }^{11,12}$ Brindle et al ${ }^{11}$ have
formulated a CVD risk prediction algorithm (QRISK) using data on $>1$ million nondiabetic patients in general practice in the United Kingdom. The QRISK (risk score using the QRESEARCH database) algorithm incorporates family history and social deprivation (in addition to the risk factors included in the Framingham risk score) and calibrates better to the UK population than the older Framingham CVD risk functions formulated by Anderson et al, ${ }^{3}$ but clinical CVD events were not formally adjudicated with a review process (as is done at Framingham). ${ }^{11} \mathrm{~A}$ formal comparison of these 2 scores on a third cohort could prove very useful.

The second risk score (ASSIGN; ASsessing cardiovascular risk using SIGN guidelines to assign potential patients to preventive treatment) from the United Kingdom was developed by investigators using $>12000$ individuals from the Scottish Heart Health Extended cohort. ${ }^{12}$ The ASSIGN score also incorporated family history and deprivation and performed marginally better than the older Framingham CVD risk functions. ${ }^{3,12}$ Additional investigations are warranted to formally compare the performance of the QRISK $^{11}$ and ASSIGN ${ }^{12}$ scores relative to the new CVD risk functions proposed here. It is conceivable that addition of other risk factors variables (such as family history or deprivation) may improve the performance of the general CVD risk function proposed in this investigation. It also is likely that risk scores for CVD developed within countries may be better calibrated for risk estimation than the Framingham general CVD risk function.

The metabolic syndrome, a much-debated multivariable risk profile, also can be used to predict CVD. This syndrome has been compared with the Framingham risk score as a predictor of CHD. ${ }^{47}$ The presence of the metabolic syndrome was found to be a significant predictor of CHD, but it was not quite as good a predictor as the Framingham risk score.

## Strengths and Limitations

The large community-based sample that is under continuous surveillance using the same standardized criteria for CVD incidence and the assessment of model performance measures such as discrimination, calibration, and exchangeability with disease-specific profiles strengthen the present investigation. However, several limitations of the present study must be acknowledged. Given the predominantly white Framingham sample, the transportability of the CVD risk function in other samples must be evaluated. Other Framingham risk functions have shown themselves to be transportable, ${ }^{19}$ at times with a recalibration. ${ }^{19-21}$ Additionally, it has been emphasized that risk scores per se do not translate to better patient outcomes unless they are used appropriately by physicians using risk communication tools and the communicated risks are well understood by the patients. ${ }^{24,48}$

## Implications

Individuals with a high global CVD risk (eg, a 10-year risk of a CVD event $>20 \%$ ) require more aggressive risk factor modification. The goal of therapy of dyslipidemia, diabetes, and hypertension should be linked to the global CVD risk. Although atherosclerotic disease-specific profiles are available, a multivariable risk formulation for global CVD made
up of standard risk factors is particularly relevant for primary prevention of atherosclerotic CVD because it is intuitive that measures taken to prevent any 1 CVD outcome can be expected to also prevent risk of the other CVD outcomes. Therefore, use of a general CVD risk score is an attractive option in office-based primary care practices. Serial assessment of global CVD risk could be used to monitor progress of patients on treatment and improvement in their multivariable risk scores.

Other risk factors not included in the general risk profile must be taken into account in evaluating risk and selecting the most efficacious treatment. These include abdominal obesity, ECG evidence of left ventricular hypertrophy, indications of insulin resistance, triglycerides, and a strong family history of premature CVD. Obesity is not included because its influence is largely attributable to its promotion of insulin resistance and its attendant CVD risk factors.

## Presentations

The CVD risk functions of Table 2 are easily programmed, for example, as an Excel spreadsheet or as the score sheets of Tables 5 through 10. This was done with the Adult Treatment Panel III cholesterol guidelines and the SCORE equations. ${ }^{5,9,12}$ In this investigation, we also present a new concept of heart age/vascular age. Here, the CVD risk of an individual is transformed to the age of a person with the same risk but all other risk factors at the normal level (nontreated systolic blood pressure of 125 mm Hg , total cholesterol of $180 \mathrm{mg} / \mathrm{dL}$, HDL of $45 \mathrm{mg} / \mathrm{dL}$, nonsmoker, nondiabetic). As seen in the example given in the Appendix, a 61 -year-old woman with risk factors above normal levels has the heart age/vascular age of a 73-year-old female with normal risk factors; similarly a 53-year-old man with risk factors has the heart age/vascular age of a man 64 years of age with normal risk factors.

## Conclusions

The present investigation presents a sex-specific multivariable risk factor algorithm that can be conveniently used to assess general CVD risk and risk of individual CVD events (coronary, cerebrovascular, and peripheral artery disease and heart failure). The estimated absolute CVD event rates predicted can be used to quantify risk and to guide preventive primary care. The validity and transportability of the proposed general CVD risk function should be evaluated in future studies.

## Appendix

## Risk Estimation From Cox Model and From

## Score Sheet

The following examples illustrate the direct application of the Cox model and the use of the score sheet to estimate CVD risk in women and men.

General formula:

$$
\hat{p}=1-S_{0}(t)^{\exp \left(\sum_{i=1}^{p} \beta_{i} X_{i}-\sum_{i=1}^{p} \beta_{i} \bar{X}_{i}\right)}
$$

where $S_{0}(t)$ is baseline survival at follow-up time $t$ (here $t=10$ years; see Table 2), $\beta i$ is the estimated regression coefficient (log hazard ratio; see Table 2), $\mathrm{X} i$ is the log-transformed

Table 11. Case 1

| Risk Factor | Value |  | Points |
| :--- | :---: | :---: | :---: |
| Age | 61 |  | 9 |
| Total cholesterol | 180 | 1 |  |
| HDL | 47 |  | 0 |
| Nontreated SBP | 124 |  | 0 |
| Treated SBP | $\ldots$ |  | 0 |
| Smoker | Yes |  | 3 |
| Diabetes | No |  | 0 |
| Point total |  | 13 |  |
| Estimate of risk, \% |  | 10.0 |  |
| Heart age/vascular age, y | 73 |  |  |

SBP indicates systolic blood pressure.
value of the $i$ th risk factor, (if continuous), $\bar{X}_{i}$ is the corresponding mean, and p denotes the number of risk factors.

Case 1—Women (baseline 10-year survival=0.95012). A 61-year-old woman not treated for high blood pressure has a total cholesterol of $180 \mathrm{mg} / \mathrm{dL}, \mathrm{HDL}$ of $47 \mathrm{mg} / \mathrm{dL}$, and systolic blood pressure of 124 mm Hg and is a current smoker but is not diabetic (see Table 11).

The risk estimate based on the Cox model is computed as follows:

$$
\begin{aligned}
& \sum_{i=1}^{p} \beta_{i} X_{i}=2.32888 * \log (\mathbf{6 1})+1.20904 * \log (\mathbf{1 8 0}) \\
& \quad-0.70833 * \log (\mathbf{4 7})+2.76157 * \log (\mathbf{1 2 4})+2.82263 * 0 \\
& +0.52873 * \mathbf{1}+0.69154 * \mathbf{0}=26.9653 . \\
& \sum_{i=1}^{p} \beta_{i} \bar{X}_{i}=2.32888 * 3.8686+1.20904 * 5.3504 \\
& \quad-0.70833 * 4.0176+2.76157 * 4.2400 \\
& \quad+2.82263 * 0.5826+0.52873 * 0.3423 \\
& \quad+0.69154 * 0.0376=26.1931 . \\
& \hat{p}=1-
\end{aligned}
$$

The points system gives a 10-year estimate of risk of $10.0 \%$. Using the Cox model directly gives $10.5 \%$. The Cox estimate for a 61 -year-old woman of normal risk is $6.7 \%$.

Case 2-Men (baseline 10-year survival=0.88936). A 53-year-old man on treatment for systolic blood pressure has a total cholesterol of $161 \mathrm{mg} / \mathrm{dL}, \mathrm{HDL}$ of $55 \mathrm{mg} / \mathrm{dL}$, and systolic blood pressure of 125 mm Hg and is diabetic but is not a current smoker (see Table 12).

The risk estimate based on the Cox model is computed as follows:

$$
\sum_{i=1}^{p} \beta_{i} X_{i}=3.06117 * \log (\mathbf{5 3})+1.12370 * \log (\mathbf{1 6 1})
$$

Table 12. Case 2

| Risk Factor | Value | Points |
| :--- | :---: | ---: |
| Age | 53 |  |
| Total cholesterol | 161 | 8 |
| HDL | 55 | 1 |
| Nontreated SBP | $\ldots$ | -1 |
| Treated SBP | 125 |  |
| Smoker | No |  |
| Diabetes | Yes |  |
| Point total |  | 13 |
| Estimate of risk, \% |  | 15.6 |
| Heart age/vascular age, y |  | 64 |

SBP indicates systolic blood pressure.

$$
\begin{aligned}
& -0.93263 * \log (\mathbf{5 5})+1.93303 * \mathbf{0}+1.99881 * \log (\mathbf{1 2 5}) \\
& +0.65451 * \mathbf{0}+0.57367 * \mathbf{1}=24.3509 . \\
& \sum_{i=1}^{p} \beta_{i} \bar{X}_{i}=3.06117 * 3.8560+1.12370 * 5.3420 \\
- & 0.93263 * 3.7686+1.93303 * 4.3544+1.99881 * 0.5019 \\
+ & 0.65451 * 0.3522+0.57367 * 0.0650=23.9802 . \\
\hat{p}= & \left.1-S_{0}(t)\right)^{\exp \left(\sum_{i=1}^{p} \beta_{i} X_{i}-\sum_{i=1}^{p} \beta_{i} \bar{X}_{i}\right)}=1-0.88936^{\exp (24.3509-23.9802)} \\
= & 0.1562 \approx 15.6 \%
\end{aligned}
$$

The points system gives a 10-year estimate of risk of $15.6 \%$. Using the Cox model directly gives $15.6 \%$. The Cox estimate for a 53 -year-old man of normal risk is $9.1 \%$.

## Sources of Funding

Support was received from the National Heart, Lung, and Blood Institute's Framingham Heart Study, National Institutes of Health, as well as from National Institutes of Health grant 5R01-NS17950. Framingham Heart Study research is supported by NIH/NHLBI contract No. N01-HC-25195. Dr Kannel was supported by the Visiting Scientist Program supported by AstraZeneca. Dr Vasan is supported in part by 2K24 HL04334 (NHLBI). Dr Cobain was supported by Unilever Research. Dr Massaro was supported in part by a grant from Unilever Research.

## Disclosures

Dr D'Agostino has served on as a consultant or on the advisory board for Sanofi and Pfizer. Dr Cobain has ownership interest in Milife.com. Dr Kannel has served on the speakers' bureau and received honoraria from BMS/Sanofi. The other authors report no conflicts.

## References

1. Cupples LA, D'Agostino RB. Section 34: some risk factors related to the annual incidence of cardiovascular disease and death in pooled repeated biennial measurements. In: Kannel WB, Wolf PA, Garrison RJ, eds. Framingham Heart Study: 30 Year Follow-Up. Bethesda, Md: US Department of Health and Human Services; 1987.
2. Jackson R, Lawes CM, Bennett DA, Milne RJ, Rodgers A. Treatment with drugs to lower blood pressure and blood cholesterol based on an individual's absolute cardiovascular risk. Lancet. 2005;365:434-441.
3. Anderson KM, Odell PM, Wilson PW, Kannel WB. Cardiovascular disease risk profiles. Am Heart J. 1991;121:293-298.
4. British Cardiac Society. JBS 2: Joint British Societies' guidelines on prevention of cardiovascular disease in clinical practice. Heart. 2005;91: v1-v52.
5. Conroy RM, Pyorala K, Fitzgerald AP, Sans S, Menotti A, De Backer G, De Bacquer D, Ducimetiere P, Jousilahti P, Keil U, Njolstad I, Oganov RG, Thomsen T, Tunstall-Pedoe H, Tverdal A, Wedel H, Whincup P, Wilhelmsen L, Graham IM, for the SCORE Project Group. Estimation of ten-year risk of fatal cardiovascular disease in Europe: the SCORE Project. Eur Heart J. 2003;24:987-1003.
6. De Backer G, Ambrosioni E, Borch-Johnsen K, Brotons C, Cifkova R, Dallongeville J, Ebrahim S, Faergeman O, Graham I, Mancia G, Manger Cats V, Orth-Gomér K, Perk J, Pyörälä K, Rodicio JL, Sans S, Sansoy V, Sechtem U, Silber S, Thomsen T, Wood D, for the Third Joint Task Force of European and Other Societies on Cardiovascular Disease Prevention in Clinical Practice. European guidelines on cardiovascular disease prevention in clinical practice: Third Joint Task Force of European and Other Societies on Cardiovascular Disease Prevention in Clinical Practice (constituted by representatives of eight societies and by invited experts). Eur Heart J. 2003;24:1601-1610.
7. Jackson R. Updated New Zealand cardiovascular disease risk-benefit prediction guide. BMJ. 2000;320:709-710.
8. Menotti A, Lanti M, Gabiti-Rosei E, Carratelli L, Cavera G, Dormi A, Gaddi A, Mancini M, Motolese M, Muiesan ML, Muntoni S, Muntoni S, Notarbartolo A, Prati P, Remiddi S, Zanchetti A. Riskard 2005: new tools for prediction of cardiovascular disease risk derived from Italian population studies. Nutr Metab Cardiovasc Dis. 2005;15:426-440.
9. Executive summary of the Third Report of The National Cholesterol Education Program (NCEP) Expert Panel on Detection, Evaluation, and Treatment of High Blood Cholesterol in Adults (Adult Treatment Panel III). JAMA. 2001;285:2486-2497.
10. Ridker PM, Buring JE, Rifai N, Cook NR. Development and validation of improved algorithms for the assessment of global cardiovascular risk in women: the Reynolds risk score. JAMA. 2007;297:611-619.
11. Hippisley-Cox J, Coupland C, Vinogradova Y, Robson J, May M, Brindle P. Derivation and validation of QRISK, a new cardiovascular disease risk score for the United Kingdom: prospective open cohort study. BMJ. 2007;335:136.
12. Woodward M, Brindle P, Tunstall-Pedoe H, for the SIGN Group on Risk Estimation. Adding social deprivation and family history to cardiovascular risk assessment: the ASSIGN score from the Scottish Heart Health Extended Cohort (SHHEC). Heart. 2007;93:172-176.
13. Anderson KM, Wilson PW, Odell PM, Kannel WB. An updated coronary risk profile: a statement for health professionals. Circulation. 1991;83: 356-362.
14. Assmann G, Cullen P, Schulte H. Simple scoring scheme for calculating the risk of acute coronary events based on the 10-year follow-up of the Prospective Cardiovascular Munster (PROCAM) Study. Circulation. 2002;105:310-315.
15. Ferrario M, Chiodini P, Chambless LE, Cesana G, Vanuzzo D, Panico S, Sega R, Pilotto L, Palmieri L, Giampaoli S. Prediction of coronary events in a low incidence population: assessing accuracy of the CUORE Cohort Study prediction equation. Int J Epidemiol. 2005;34:413-421.
16. Wilson PW, D'Agostino RB, Levy D, Belanger AM, Silbershatz H, Kannel WB. Prediction of coronary heart disease using risk factor categories. Circulation. 1998;97:1837-1847.
17. Zhang XF, Attia J, D'Este C, Yu XH, Wu XG. A risk score predicted coronary heart disease and stroke in a Chinese cohort. J Clin Epidemiol. 2005;58:951-958.
18. Wolf PA, D'Agostino RB, Belanger AJ, Kannel WB. Probability of stroke: a risk profile from the Framingham study. Stroke. 1991;22: 312-318.
19. Murabito JM, D'Agostino RB, Silbershatz H, Wilson WF. Intermittent claudication: a risk profile from the Framingham Heart Study. Circulation. 1997;96:44-49.
20. Kannel WB, D'Agostino RB, Silbershatz H, Belanger AJ, Wilson PW, Levy D. Profile for estimating risk of heart failure. Arch Intern Med. 1999;159:1197-1204.
21. D'Agostino S, Grundy S, Sullivan LM, Wilson P, for the CHD Risk Prediction Group. Validation of the Framingham coronary heart disease prediction scores: results of a multiple ethnic groups investigation. JAMA. 2001;286:180-187.
22. Liu J, Hong Y, D’Agostino RB Sr, Wu Z, Wang W, Sun J, Wilson PW, Kannel WB, Zhao D. Predictive value for the Chinese population of the Framingham CHD risk assessment tool compared with the Chinese MultiProvincial Cohort Study. JAMA. 2004;291:2591-2599.
23. Marrugat J, D’Agostino R, Sullivan L, Elosua R, Wilson P, Ordovas J, Solanas P, Cordon F, Ramos R, Sala J, Masia R, Kannel WB. An adaptation of the Framingham coronary heart disease risk function to European Mediterranean areas. J Epidemiol Community Health. 2003;57: 634-638.
24. Beswick A, Brindle P. Risk scoring in the assessment of cardiovascular risk. Curr Opin Lipidol. 2006;17:375-386.
25. Dawber TR, Meadors GF, Moore FE. Epidemiologic approaches to heart disease: the Framingham study. Am J Public Health. 1951;41:279-286.
26. Kannel WB, Feinleib M, McNamara PM, Garrison RJ, Castelli WP. An investigation of coronary heart disease in families: the Framingham offspring study. Am J Epidemiol. 1979;110:281-290.
27. Abbott R, McGee D. The probability of developing certain cardiovascular diseases in eight years at specified values of some characteristics. In: The Framingham Study: An Epidemiological Investigation of Cardiovascular Disease. Bethesda, Md: National Heart, Lung, and Blood Institute; 1987:sec 37.
28. Cox DR. Regression models and life tables. J Royal Stat Soc. 1972; 34(series B):187-220.
29. D'Agostino R, Nam BH. Evaluation of the performance of survival analysis models: discrimination and calibration measures. In: Handbook of Statistics. Amsterdam, The Netherlands: Elsevier; 2004:1-25.
30. Pencina MJ, D'Agostino RB. Overall C as a measure of discrimination in survival analysis: model specific population value and confidence interval estimation. Stat Med. 2004;23:2109-2123.
31. Harrell FE Jr, Lee KL, Mark DB. Multivariable prognostic models: issues in developing models, evaluating assumptions and adequacy, and measuring and reducing errors. Stat Med. 1996;15:361-387.
32. Antolini L, Nam BH, Agostino RB. Inference on correlated discrimination measures in survival analysis: a nonparametric approach. Commun Stat Theory Methods. 2004;33:2117-2135.
33. Pencina MJ, D'Agostino RB Sr, D'Agostino RB Jr, Vasan RS. Evaluating the added predictive ability of a new marker: from area under the ROC curve to reclassification and beyond. Stat Med. 2008;27:157-172.
34. Sullivan LM, Massaro JM, D'Agostino RB Sr. Presentation of multivariate data for clinical use: the Framingham study risk score functions. Stat Med. 2004;23:1631-1660.
35. Rosamond W, Flegal K, Friday G, Furie K, Go A, Greenlund K, Haase N, Ho M, Howard V, Kissela B, Kittner S, Lloyd-Jones D, McDermott M, Meigs J, Moy C, Nichol G, O’Donnell CJ, Roger V, Rumsfeld J, Sorlie P, Steinberger J, Thom T, Wasserthiel-Smoller S, Hong Y. Heart disease and stroke statistics: 2007 update: a report from the American Heart Association Statistics Committee and Stroke Statistics Subcommittee. Circulation. 2007;115:e69-e171.
36. Lopez AD, Mathers CD, Ezzati M, Jamison DT, Murray CJ. Global and regional burden of disease and risk factors, 2001: systematic analysis of population health data. Lancet. 367:1747-1757.
37. Lloyd-Jones DM, Leip EP, Larson MG, D'Agostino RB, Beiser A, Wilson PWF, Wolf PA, Levy D. Prediction of lifetime risk for cardiovascular disease by risk factor burden at 50 years of age. Circulation. 2006;113:791-798.
38. Furberg CD, Hennekens CH, Hulley SB, Manolio T, Psaty BM, Whelton PK. 27th Bethesda Conference: matching the intensity of risk factor management with the hazard for coronary disease events: Task Force 2: clinical epidemiology: the conceptual basis for interpreting risk factors. J Am Coll Cardiol. 1996;27:976-978.
39. Grundy SM. Primary prevention of coronary heart disease: integrating risk assessment with intervention. Circulation. 1999;100:988-998.
40. Kannel WB, D'Agostino RB, Sullivan L, Wilson PW. Concept and usefulness of cardiovascular risk profiles. Am Heart J. 2004;148:16-26.
41. Yusuf S, Hawken S, Unpuu S, Dans T, Avezum A, Lanas F, McQueen M, Budaj A, Pais P, Varigos J. Effect of potentially modifiable risk factors associated with myocardial infarction in 52 countries (the INTERHEART study): case-control study. Lancet. 364:937-952.
42. Califf RM, Armstrong PW, Carver JR, D'Agostino RB, Strauss WE. 27th Bethesda Conference: matching the intensity of risk factor management with the hazard for coronary disease events: Task Force 5: stratification of patients into high, medium and low risk subgroups for purposes of risk factor management. J Am Coll Cardiol. 1996;27:1007-1019.
43. Forrester JS, Merz CN, Bush TL, Cohn JN, Hunninghake DB, Parthasarathy S, Superko HR. 27th Bethesda Conference: matching the intensity of risk factor management with the hazard for coronary disease events: Task Force 4: efficacy of risk factor management. J Am Coll Cardiol. 1996;27:991-1006.
44. Goldman L, Garber AM, Grover SA, Hlatky MA. 27th Bethesda Conference: matching the intensity of risk factor management with the hazard for coronary disease events: Task Force 6: cost effectiveness of assessment and management of risk factors. J Am Coll Cardiol. 1996;27: 1020-1030.
45. Chang M, Hahn RA, Teutsch SM, Hutwagner LC. Multiple risk factors and population attributable risk for ischemic heart disease mortality in the United States, 1971-1992. J Clin Epidemiol. 2001;54:634-644.
46. Kannel WB, McGee D, Gordon T. A general cardiovascular risk profile: the Framingham study. Am J Cardiol. 1976;38:46-51.
47. Wannamethee SG, Shaper AG, Lennon L, Morris RW. Metabolic syndrome vs Framingham risk score for prediction of coronary heart disease, stroke, and type 2 diabetes mellitus. Arch Intern Med. 2005;165: 2644-2650.
48. Brindle P, Beswick A, Fahey T, Ebrahim S. Accuracy and impact of risk assessment in the primary prevention of cardiovascular disease: a systematic review. Heart. 2006;92:1752-1759.

## CLINICAL PERSPECTIVE

Cardiovascular disease (CVD) remains the leading cause of mortality and morbidity. A plethora of effective drugs for the major risk factors of CVD such as blood pressure, high cholesterol, and diabetes control exists. Prevention strategies assessing the risk of CVD and identifying the risk factors associated with the risk have become major approaches for reducing CVD and its unfavorable consequences. Presently, many of these strategies involve focusing on a component of CVD such as hard coronary disease consisting of myocardial infarction and coronary death, assessing the risk by mathematical risk functions or scoring functions, and designing treatment (behavioral and/or drug) according to the level of risk. Our belief is that, especially in the primary care setting, CVD risk should not be directed only to a component of CVD such as coronary disease or stroke but rather to all manifestation of CVD. In particular, individuals with high overall CVD risk require aggressive risk factor modification. In the present article, we present simple sex-specific risk functions that assess the 10 -year risk (probability) of developing overall CVD. These functions involve inputs of blood pressure, cholesterol levels, diabetes, and smoking. Thus, not only is the overall risk quantified, but the source of the risk can be identified for treatment. Simple scoring sheets are presented that make the CVD functions immediately usable. Finally, straightforward adjustments to the functions can be used also to assess the risk for specific components of CVD.

## Correction

In the article by D'Agostino, et al, "General Cardiovascular Risk Profile for Use in Primary Care: The Framingham Heart Study," which appeared in the February 12, 2008, issue of the journal (Circulation. 2008;117:743-753), the following grant should have been listed as a funding source: National Institutes of Health 5R01-NS17950.

The error has been corrected in the current online version of the article.

The authors regret this omission.
DOI: 10.1161/CIRCULATIONAHA.108.190154

# General Cardiovascular Risk Profile for Use in Primary Care: The Framingham Heart Study <br> Ralph B. D'Agostino, Sr, Ramachandran S. Vasan, Michael J. Pencina, Philip A. Wolf, Mark 

 Cobain, Joseph M. Massaro and William B. KannelCirculation. 2008;117:743-753; originally published online January 22, 2008;
doi: 10.1161/CIRCULATIONAHA.107.699579
Circulation is published by the American Heart Association, 7272 Greenville Avenue, Dallas, TX 75231
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Print ISSN: 0009-7322. Online ISSN: 1524-4539

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[^0]:    Received February 27, 2007; accepted November 30, 2007.
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    Guest Editor for this article was Eric B. Rimm, ScD.
    The online Data Supplement can be found with this article at http://circ.ahajournals.org/cgi/content/full/CIRCULATIONAHA.107. 699579/DC1.

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